Trade and Unemployment in Germany: An Empirical Exploration and Some Theories

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Abstract—This paper investigates the link between unemployment and external trade in Germany, using data on unemployment, international trade and economic activity for the 16 German federal states (Länder). With panel data econometrics we show that international trade as measured by state-by-state trade shares has a significantly negative impact on state-by-state unemployment rates. We investigate to what extent this reflects mercantilist tendencies embedded in the political economy of the Federal Republic but we find no clear evidence in favor of simple mercantilism. We suggest three alternative explanations, which are linked to the structural rigidities of the German labor market, government involvement in the education and training system and certain labor market policies which provide government-subsidized unemployment insurance benefitting (mainly) export industries.

Keywords—JEL; F14; F15; O50; International Trade; Unemployment; Panel Study

I. INTRODUCTION

The connection between external trade and unemployment is often addressed by politicians but seldom by economists. Among the latter, the idea of promoting exports in order to reduce unemployment counts as a relapse into the dark ages of mercantilism. As a matter of course, mainstream economics orthogonalizes external trade and unemployment. By assumption, the real wage clears the labor market and the terms of trade equilibrate external trade. Our paper goes against this prevailing attitude and investigates the link between unemployment and external trade. The focus of our investigation is on the German Länder (federal states) and their performance in unemployment and external trade since German unification in 1990. Western states and Eastern states exhibited significant differences in these areas. Eastern states were hit by the collapse of Soviet bloc trade in the early 90s while Western states were hit by the collapse of world trade in the wake of the recent financial crisis. Using panel data econometrics we find that international trade measured by state-specific exports has a significant negative impact on state-specific unemployment. We test to what extent this reflects mercantilist tendencies embedded in the German political economy, where mercantilism is a government policy aiming for trade surpluses in order to maintain and promote domestic employment. We find no clear evidence in favor of simple mercantilism. Potential reasons for our observations are the existence of minimum wages implicated in the German welfare state, rigidities in the education system and the institutional details of the unemployment insurance system. Regarding the latter, we highlight in particular the workings of the so-called short work system (Kurzarbeit), which can be viewed as a government-provided insurance scheme with benefits accruing mainly to export industries.

The remainder of our paper is organized as follows. In the following section II we present the data and discuss the methodology that we use to establish some of our series. We also present stylized facts of the empirical link between exports and unemployment in Germany. After that, in section III, we investigate whether the observed stylized facts can be attributed to "mercantilism". Section IV discusses the factors that might explain the observed link between external trade and unemployment.

II. DATA SOURCES AND STYLIZED FACTS

Our study is based on a set of annual data on unemployment, economic activity and external trade in Germany from 1990 the re-unification of the country to the present. The data, their acronyms and their sources are summarized in Table I. The special flavor of our study derives from the fact that our data are regionally disaggregated, i.e. they provide information about the German federal states (Länder). Since unification in 1990, Germany has been divided into 16 federal states: 5 states from socialist former East Germany, 10 states from former West Germany and one state - the city state of Berlin - in a hybrid position in so far as it is the combination of former West and East (i.e. socialist) Berlin. For each of these states our data set contains nominal and real GDP, merchandise exports and imports, as well as several unemployment indicators. The source of the data on GDP and on trade is the German federal statistical office, the source of all unemployment data is the German federal employment agency (Arbeitsagentur). Unemployment data are on one hand official state-by-state unemployment rates, on the other hand several measures of short work, known as Kurzarbeit. Under the rules of short work, firms can reduce hours by 25 to 100 percent while keeping their employees on the payroll. Around 65 percent of the employees’ reduction in wages is then paid by the government. The firms can thus lower their wage payments by the respective percentage while the government compensates employees for the shortfall in earnings. The so-called Kurzarbeit Null (short work zero) basically amounts to a temporary layoff where the government continues to pay employees’ wages (though not in full). Available data on short work are grouped by the degree of the reduction in hours, namely 25, 50, 75 and 100 percent (the latter corresponding to Kurzarbeit Null). As short work is officially not considered unemployment, these data are not included in the official unemployment rates. In order to correct this, we use a hybrid measure of unemployment by adding short work, appropriately weighted by the respective degree of reduction in hours, to official unemployment, and expressing the resulting numbers of percent of the labor force.
Our first step is to investigate the data set by establishing stylized facts concerning unemployment, growth and exports among the German federal states. Since these are embedded in a single legal and institutional framework, the data set lends itself for applying panel data econometrics.

We begin by running a simple OLS regression of state-specific unemployment rates on state-specific export-to-GDP ratios using a pooled data set of all federal states and time periods:

$$U_t = \beta_0 + \beta_1 EX_t + \nu_t$$  \hspace{1cm} (1)

The index \(i = 1, \ldots, 16\) identifies the different federal states. The time index \(t = 1, 2, 3, \ldots, T\) identifies the annual observations. The \(\nu_t\) are iid error terms. The results of the regression are found in the first column of Table II. The regression exhibits a highly significant negative slope coefficient, seemingly indicating a strong improvement in unemployment resulting from an increase in exports. However, this is a classic example of heterogeneity bias. The negative coefficient largely reflects the differences between the old and the new federal states. The former has high unemployment rates and low export-to-GDP ratios during the sample period (resulting, respectively, from economic restructuring and from the collapse of trade within the former Soviet bloc); the latter has relatively lower unemployment rates and relatively higher export-to-GDP ratios. Thus the negative coefficient simply indicates that Western states have had lower unemployment rates and stronger exports than the Eastern states during the sample period.

Our next regression equation handles this heterogeneity bias by introducing fixed effects:

$$U_t = \delta \beta_0 + \beta_1 EX_t + \mu_i + \nu_t$$  \hspace{1cm} (2)

The coefficients \(\mu_i\), for \(i = 1, \ldots, 16\), reflect the state-specific fixed effects. The rest of the notation is as above. Column (2) in Table II shows the results of an LSDV regression of this equation. The results show that once heterogeneity is factored in, the expected sign of the real growth rate is weak, and second, by using a lag we can avoid endogeneity problems on the right-hand side. Column (3) in Table II shows the results of the within estimators applied to this equation. There is strong positive persistence in the unemployment rate, a significant negative impact of real growth on unemployment and, finally, a significant negative impact of exports on unemployment.

In order to check for robustness we introduce alternative estimation methods. For autoregressive panel data we models ordinary LSDV estimation or within estimation give inconsistent results for finite T (see e.g. [16]). The econometric literature offers several consistent estimators, using instruments for the lagged dependent as in [1], [2] and extensions like [4]. All these consistent estimators were mainly intended for panels with large N and (relatively) small T. In our panel, T is approximately equal to N. Kiviet [12] proposed a correction for the bias in the LSDV estimator. Further results for correcting the bias were presented in [13] and evaluated in [7]. Bruno ([5], [6]) showed that this correction performs well as \(N < T\) and also in unbalanced panels.

In order to check whether our results are susceptible to bias due to the presence of the lagged endogenous on the right-hand side, we compare the results from the within estimator with those obtained with the bias-corrected estimator proposed by [13] in column (4) and the Arellano-Bond estimator in column (5). All three methods yield roughly the same results.

So far our regressions have used the official unemployment rate as dependent variable. Now we take into account the impact of active unemployment policy, in particular of the short work scheme which is an important part of German labor market policy. We begin by setting up a regression equation structurally similar to the previous one, albeit with short work replacing unemployment:

$$SW_t = \delta SW_{t-1} + \beta_0 RG_{t-1} + \beta_1 EX_{t-1} + \mu_i + \nu_t$$  \hspace{1cm} (4)

The results – again from the three estimators used previously – are reported in Table III. There are significant differences from the results of the regressions using the unemployment rate as endogenous variable. Persistence of short work – as indicated by the coefficient of the first-order lag – is much weaker, as is the negative impact of the real growth rate. The impact of the export-to-GDP ratio, on the other hand, is high-

<table>
<thead>
<tr>
<th>Description</th>
<th>Format</th>
<th>Source</th>
<th>Obs.</th>
</tr>
</thead>
<tbody>
<tr>
<td>RG</td>
<td>% p.a.</td>
<td>Regional Accounts (2010)</td>
<td>304</td>
</tr>
<tr>
<td>EX</td>
<td>% of GDP</td>
<td>Federal Statistical Office (2010)</td>
<td>304</td>
</tr>
<tr>
<td>IM</td>
<td>% of GDP</td>
<td>Federal Statistical Office (2010)</td>
<td>304</td>
</tr>
<tr>
<td>ET</td>
<td>% of GDP</td>
<td>EX + IM</td>
<td>304</td>
</tr>
<tr>
<td>TB</td>
<td>% of GDP</td>
<td>EX + IM</td>
<td>304</td>
</tr>
<tr>
<td>U</td>
<td>% of labor force</td>
<td>Federal Employment Agency (2010)</td>
<td>304</td>
</tr>
<tr>
<td>SW</td>
<td>% of labor force</td>
<td>Federal Employment Agency (07/2010)</td>
<td>272</td>
</tr>
<tr>
<td>AU</td>
<td>% of labor force</td>
<td>Federal Employment Agency (07/2010 &amp; 2010)</td>
<td>272</td>
</tr>
</tbody>
</table>

1 They show that also a simplified version of [13] correction term is sufficient.
ly significant and much stronger in relative terms. This suggests a particular link between export fluctuations and the activation of short work compensation. This, in turn, suggests that better results can be obtained by combining the official unemployment rate with the short work rate into one single compound adjusted unemployment measure. As described above, we do this with our adjusted unemployment rate, AU. Our final step in this section is then to use this adjusted unemployment rate in a regression on growth and the exports-to-GDP ratio.

### TABLE II

**BASIC REGRESSIONS WITH UNEMPLOYMENT RATE AS THE DEPENDENT VARIABLE**

<table>
<thead>
<tr>
<th>Methodology</th>
<th>OLS</th>
<th>LSDV</th>
<th>Within</th>
<th>Kiviet</th>
<th>Arellano-Bond</th>
</tr>
</thead>
<tbody>
<tr>
<td>EX</td>
<td>-0.2207</td>
<td>***</td>
<td>-0.0357</td>
<td>***</td>
<td>-0.1020 ***</td>
</tr>
<tr>
<td>U(-1)</td>
<td>0.7498</td>
<td>***</td>
<td>0.8052</td>
<td>***</td>
<td>0.7285 ***</td>
</tr>
<tr>
<td>RG(-1)</td>
<td>-0.1309</td>
<td>***</td>
<td>-0.1210</td>
<td>***</td>
<td>-0.1415 ***</td>
</tr>
<tr>
<td>cons</td>
<td>17.6662</td>
<td>***</td>
<td>13.6769</td>
<td>***</td>
<td>5.6254 ***</td>
</tr>
<tr>
<td>N</td>
<td>272</td>
<td>272</td>
<td>272</td>
<td>256</td>
<td>256</td>
</tr>
</tbody>
</table>

Note: The Kiviet estimation procedure applied here gives no constant in column (4). It is adjusted to the fixed effects. The N differs due to differencing. Asterisks indicate p-values of the coefficients, with p<0.05 (*), p<0.01 (**) and p<0.001 (**).

### TABLE III

**REGRESSION WITH SHORT WORK RATE AND ADJUSTED UNEMPLOYMENT RATE AS THE DEPENDENT VARIABLE**

<table>
<thead>
<tr>
<th>Methodology</th>
<th>Kiviet</th>
<th>Arellano-Bond</th>
<th>Within</th>
<th>Kiviet</th>
<th>Arellano-Bond</th>
</tr>
</thead>
<tbody>
<tr>
<td>EX</td>
<td>-0.0467</td>
<td>***</td>
<td>-0.0467</td>
<td>***</td>
<td>-0.0972 ***</td>
</tr>
<tr>
<td>SW(-1)</td>
<td>-0.0519</td>
<td>-0.1387 *</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>AU(-1)</td>
<td>0.6529</td>
<td>***</td>
<td>0.7065</td>
<td>***</td>
<td>0.6217 ***</td>
</tr>
<tr>
<td>RG(-1)</td>
<td>0.0473</td>
<td>**</td>
<td>-0.1295</td>
<td>***</td>
<td>-0.1197 ***</td>
</tr>
<tr>
<td>cons</td>
<td>1.8104</td>
<td>***</td>
<td>6.9942</td>
<td>***</td>
<td>7.5078 ***</td>
</tr>
<tr>
<td>N</td>
<td>271</td>
<td>255</td>
<td>272</td>
<td>272</td>
<td>256</td>
</tr>
</tbody>
</table>

\[ AU = \delta \cdot U + \beta \cdot RG + \beta \cdot EX + \mu + \nu \]  

The results (cf. Table III) are similar to the regression results for the simple unemployment rate although there are some differences in details. The autoregressive component is now somewhat weaker while the negative impact of the export-to-GDP ratio slightly stronger in absolute terms. We treat this as our preferred equation with which we will continue to work in the following section.

### III. IS THERE EVIDENCE OF "MERCANTILISM"?

The stylized facts established in the previous section suggest that export performance contributes to reducing unemployment in Germany. How can these results be interpreted? The first interpretation – on which we focus in this section – is that this reflects a sort of “mercantilism” embedded in the political economy of Germany. This point is succinctly expressed by [8] who speaks of “German mercantilist model”, stating that “…in the German case, the national mystique of a trade surplus may have had a role in disciplining the labour market and at the same time assuring profits.” He suggests that “successful exports and depressed imports (due to restrained domestic consumption) lead to a trade surplus that becomes the icon of the success of the model and the morning star of social and economic policy, in particular labour market discipline.” (ibid. p. 8) This eloquent statement provides a convenient anchoring point for the following Hypothesis:

German political economy is characterized by embedded mercantilism, i.e. a policy stance geared at ensuring low wages and trade surpluses. The observed negative link between exports and unemployment is an evidence of this. In addition one should observe that (1) a rise in imports leads to a worsening of unemployment and (2) an improvement in the trade balance leads to an improvement in unemployment.

If imports and the trade balance are added to right-hand side of our unemployment regression then one should expect a positive coefficient for imports and a negative coefficient for the trade balance.

In Table IV columns (11) to (14) we integrate state-by-state the import-to-GDP ratios into our preferred regression
The correlations reported in Table V show that the correlation coefficient between total trade and the trade balance is significantly different from zero but with a value of approximately 0.11 which is small enough that the two regressors are not multicollinear. Adding to the trade balance our preferred regression equation has the further advantage that we can test part (2) of our above hypothesis on German mercantilism. Our regression equation now looks like this:

\[
AU = \alpha + \beta_1 AU_{t-1} + \beta_2 RG_{t-1} + \beta_3 TT_{t-1} + \beta_4 TB_{t-1} + \mu_t + \nu_t
\]

(8)

The results are reported in Table VI (columns (15) and (16)). The impact of total trade on unemployment is negative, with a coefficient approximately half of the coefficient of exports reported in the previous section. This is not surprising, given that exports and imports are strongly positively correlated and of similar size (i.e. total trade is approximately twice the size of exports and move roughly in parallel with it). The coefficient of the trade balance is negative but not significant. Thus part (2) of the hypothesis on mercantilism is not confirmed.

In a final step we can use these coefficients to derive a consistent test for the size of the implied coefficient of imports. Using the definitions \( TT = EX + IM \) and \( TB = EX - IM \) we have:

\[
\beta_1 TT_{t-1} + \beta_3 TB_{t-1} = \\
(\beta_1 + \beta_3)EX_{t-1} + (\beta_3 - \beta_3)IM_{t-1}
\]

(9)

The definition of the implied coefficient \( \beta_5 = \beta_1 - \beta_3 \) is used in Table VII for another test of part (1) of our hypothesis on mercantilism.

### Table V

<table>
<thead>
<tr>
<th></th>
<th>IM</th>
<th>TT</th>
<th>TB</th>
</tr>
</thead>
<tbody>
<tr>
<td>EX</td>
<td>0.8610</td>
<td>0.9666</td>
<td>0.3602</td>
</tr>
<tr>
<td>IM</td>
<td>0.9626</td>
<td>-0.1643</td>
<td></td>
</tr>
<tr>
<td>TT</td>
<td>0.1091</td>
<td></td>
<td></td>
</tr>
</tbody>
</table>

No information is lost in moving from the right-hand side with exports and imports to the left-hand side with total trade and the trade balance. By inverting this matrix equation one can express exports and imports as linear transformations of total trade and the trade balance:

\[
\begin{pmatrix} EX \\ IM \end{pmatrix} = \begin{pmatrix} 1 & 1 \\ 1 & -1 \end{pmatrix} \begin{pmatrix} TT \\ TB \end{pmatrix}
\]

(7)

From the point of view of economic theory, it is not clear which one of the two pairs – (EX, IM) or (TT, TB) – is the "fundamental" one. We are thus free to use the statistically better behaved pair of data series – (TT, TB) – in our regression equation.

The correlations reported in Table V show that the correlation coefficient between total trade and the trade balance is significantly different from zero but with a value of approximately 0.11 which is small enough that the two regressors are not multicollinear. Adding to the trade balance our preferred regression equation has the further advantage that we can test part (2) of our above hypothesis on German mercantilism. Our regression equation now looks like this:

\[
AU = \alpha + \beta_1 AU_{t-1} + \beta_2 RG_{t-1} + \beta_3 TT_{t-1} + \beta_4 TB_{t-1} + \mu_t + \nu_t
\]

(8)

The results are reported in Table VI (columns (15) and (16)). The impact of total trade on unemployment is negative, with a coefficient approximately half of the coefficient of exports reported in the previous section. This is not surprising, given that exports and imports are strongly positively correlated and of similar size (i.e. total trade is approximately twice the size of exports and move roughly in parallel with it). The coefficient of the trade balance is negative but not significant. Thus part (2) of the hypothesis on mercantilism is not confirmed.

In a final step we can use these coefficients to derive a consistent test for the size of the implied coefficient of imports. Using the definitions \( TT = EX + IM \) and \( TB = EX - IM \) we have:

\[
\beta_1 TT_{t-1} + \beta_3 TB_{t-1} = \\
(\beta_1 + \beta_3)EX_{t-1} + (\beta_3 - \beta_3)IM_{t-1}
\]

(9)

The definition of the implied coefficient \( \beta_5 = \beta_1 - \beta_3 \) is used in Table VII for another test of part (1) of our hypothesis on mercantilism.

---

2 The correlations are computed as within correlations, i.e. after removal of state-specific averages. This is done by left-multiplying each series with the matrix \( Q = I - P \), where \( I \) is the identity matrix and \( P \) is a matrix which computes state-specific averages as explained in [3], p. 14f.
TABLE VI
REGRESSION RESULTS WITH ADJUSTED UNEMPLOYMENT RATE AS THE DEPENDENT VARIABLE WITH RESPECT TO EXPORT/IMPORT CHARACTERISTIC

<table>
<thead>
<tr>
<th></th>
<th>(15)</th>
<th>(16)</th>
</tr>
</thead>
<tbody>
<tr>
<td>Methodology</td>
<td>Kiviet</td>
<td>Arellano-Bond</td>
</tr>
<tr>
<td>TT</td>
<td>-0.0511 ***</td>
<td>-0.0547 ***</td>
</tr>
<tr>
<td>TB</td>
<td>-0.0154</td>
<td>-0.0068</td>
</tr>
<tr>
<td>AU(-1)</td>
<td>0.7132 ***</td>
<td>0.6326 ***</td>
</tr>
<tr>
<td>RG(-1)</td>
<td>-0.1170 ***</td>
<td>-0.1377 ***</td>
</tr>
<tr>
<td>cons</td>
<td>7.5566 ***</td>
<td></td>
</tr>
<tr>
<td>N</td>
<td>272</td>
<td>256</td>
</tr>
</tbody>
</table>

Mercantilism would mean that the implied coefficient of imports is positive: \( \beta_{\text{IM}} > 0 \).

The results show that the probability of a parameter value in the “mercantilist range” (i.e. a positive impact of imports on unemployment) is generally low but depends on the estimator used.\(^3\) For the Kiviet estimator and the Arellano-Bond estimator mercantilism can be rejected at the 10 percent and 5 percent level, respectively. If we combine this with the result that the impact of the trade balance on unemployment is not significantly different from zero, we can conclude that there is no convincing evidence in favor of mercantilism. We can thus reject the hypothesis formulated at the beginning of this section.

TABLE VII
TEST OF IMPLIED IMPORT COEFFICIENT

<table>
<thead>
<tr>
<th></th>
<th>Kiviet</th>
<th>Arellano-Bond</th>
</tr>
</thead>
<tbody>
<tr>
<td>( \beta_{\text{IM}} )</td>
<td>-0.0357</td>
<td>-0.0479</td>
</tr>
<tr>
<td>St. dev.</td>
<td>0.0244</td>
<td>0.0242</td>
</tr>
<tr>
<td>Prob.: ( \beta_{\text{IM}} \geq 0 )</td>
<td>0.0714</td>
<td>0.0238</td>
</tr>
</tbody>
</table>

IV. EXPORTS AND UNEMPLOYMENT: BEYOND MERCANTILISM

Our results so far suggest (i) that German exports contribute significantly to reducing unemployment in the German federal states but (ii) there is no convincing evidence that reflects the case of straightforward mercantilist policies in the sense that it is the trade balance which produces this unemployment-reducing effect. What then are the reasons for the observed stylized facts? In this section we discuss some possible answers to this question.

Our counterfactual starting point is the classical foreign trade theory which typically assumes full employment. More precisely, classical foreign trade theory assumes a perfectly flexible labor market with (in particular) perfect intersectoral mobility of labor. This permits the real wage to clear the labor market and wipes out the correlation between external trade and employment. The key then to understanding why such a correlation exists in the data lies in removing this assumption and in identifying imperfections and rigidities in the labor market and other parts of the economy. Imperfections and rigidities bearing on our problem are not hard to be identified in Germany. We will highlight the following three points:

- Minimum wages
- Rigidities in the education system and generally in human capital formation
- Government policies aimed at correcting market failures and providing implicit subsidies to export-oriented industries

These three points can conveniently be linked to well-established theories of international trade: the first to Ricardian trade theories, the second to Heckscher-Ohlin type trade theories and the third to the new and “new new” trade theories developed e.g. by [14] and [15].

The first point provides a Ricardian explanation for the positive (i.e. improving) impact of exports on unemployment. Figure 1 illustrates the argument. One of Ricardo’s key conclusions in presenting his theory of comparative advantage was that free trade would lead to a rise in real wages. The figure represents an expansion of exports (due, e.g. to a rise in world market demand) as an upward shift of the labor demand curve. With labor supply inelastic and fixed, the result is a rise in the real wage from \( w_0 \) to \( w_1 \). In the presence of the minimum wage above the initial market-clearing wage \( w_0 \), the result is not a wage increase but a reduction in unemployment from \( u_0 \) to \( u_1 \). Although Germany has no nation-wide minimum wage, minimum wage laws affect certain sectors (among them several service sectors). This is complemented by the rules of the welfare state which introduce a minimum income guaranteed by the government. Moreover, binding wage agreements with strong unions in particular in export-oriented industries keep real wages high and downwardly rigid. There is thus a reason to believe that minimum wages provide one of the reasons for the observed improving impact of external trade on unemployment in Germany.

![Fig. 1 Exports and Unemployment with Minimum Wages](image)

Our second point provides a Heckscher-Ohlin type explanation for the observed link between exports and unemployment. Its basic ingredients are rigidities in the education system and in human capital formation. This somewhat fuzzy description summarizes a multitude of very concrete government policies. The point of departure is the recognition that government-provided services furnish important factor inputs to export-oriented industries (as well as for other industries). Some of these are public goods, some are private goods. Higher education and vocational training have both a public and a private good component. In Germany (as generally in Europe) government is deeply involved in higher education and in vocational training. The supply of these services is not driven by pure profit motives. Public debate about funding for

\(^3\) The probabilities in Table VII are calculated under the assumption of asymptotic normality.
unemployment. Kurzarbeit, the government obtains a reduction in long-term short-term unemployment compensation in the guise of

- Exports raise wages which, in the presence of minimum wages, reduces unemployment;
- Structural rigidities in labor markets linked, in particular, to heavy government involvement in higher education and in vocational training contribute to reducing mobility between different segments of the labor markets;
- The government-sponsored short work scheme ensures employees and employers against export volatility and thereby reduces persistence in unemployment.

Factors (1) and (2) are more or less familiar with the long-standing debate about the rigidities of European economies (Euroscerosis). The novel aspect is the insurance aspect of the short work scheme.

V. SUMMARY AND CONCLUSION

In this paper we have investigated the link between exports and unemployment in Germany. Going through a number of different regressions, equations we have arrived at a preferred equation which had demonstrated that exports (or, equivalently, total trade, i.e. the sum of exports and imports) have a significantly negative impact on unemployment rates. From a practical point of view most economists would probably feel that this link between exports and unemployment is due to the impact of exports on growth, export demand being an important driver of growth at least within a shorter horizon. One could call this the "mainstream model": causality runs from exports to growth and from growth to employment. The mainstream model rests on the (explicit or implicit) assumption of a flexible labor market. The structure of aggregate demand does no matter with employment, only its overall size. Our results contest the mainstream view. No labor market is perfectly flexible, but the German labor market in particular has been known for its rigidities and its weak intersectional mobility. Structural change away from industry towards services has been slower in Germany than in other industrial countries. Government-backed institutions such as "short work" mitigate the pressure towards structural change during slowdowns and crises and thus serve to maintain existing structures. In such an environment it can no longer be taken for granted that it is only GDP growth as a whole which is causal for unemployment. Rather, the structure of demand begins to matter with unemployment.

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